

# Section 10.2 - More Detail About Simple Linear Regression

Statistics 104

Autumn 2004



# Analysis of Variance for Regression

```
. regress cityfuel Weight
```

Source	SS	df	MS	Number of obs = 93		
Model	68.8245208	1	68.8245208	F( 1, 91)	=	374.31
Residual	16.7322059	91	.183870394	Prob > F	=	0.0000
Total	85.5567267	92	.929964421	R-squared	=	0.8044
				Adj R-squared	=	0.8023
				Root MSE	=	.4288

cityfuel	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
Weight	.0014662	.0000758	19.35	0.000	.0013157	.0016168
_cons	.1936668	.2370884	0.82	0.416	-.2772802	.6646138

The Analysis of Variance (ANOVA) Table is an alternative approach to examining a regression model.

The idea behind it is based on

$$\text{DATA} = \text{FIT} + \text{RESIDUAL}$$

The variance in the data  $y$  is expressed by the deviations

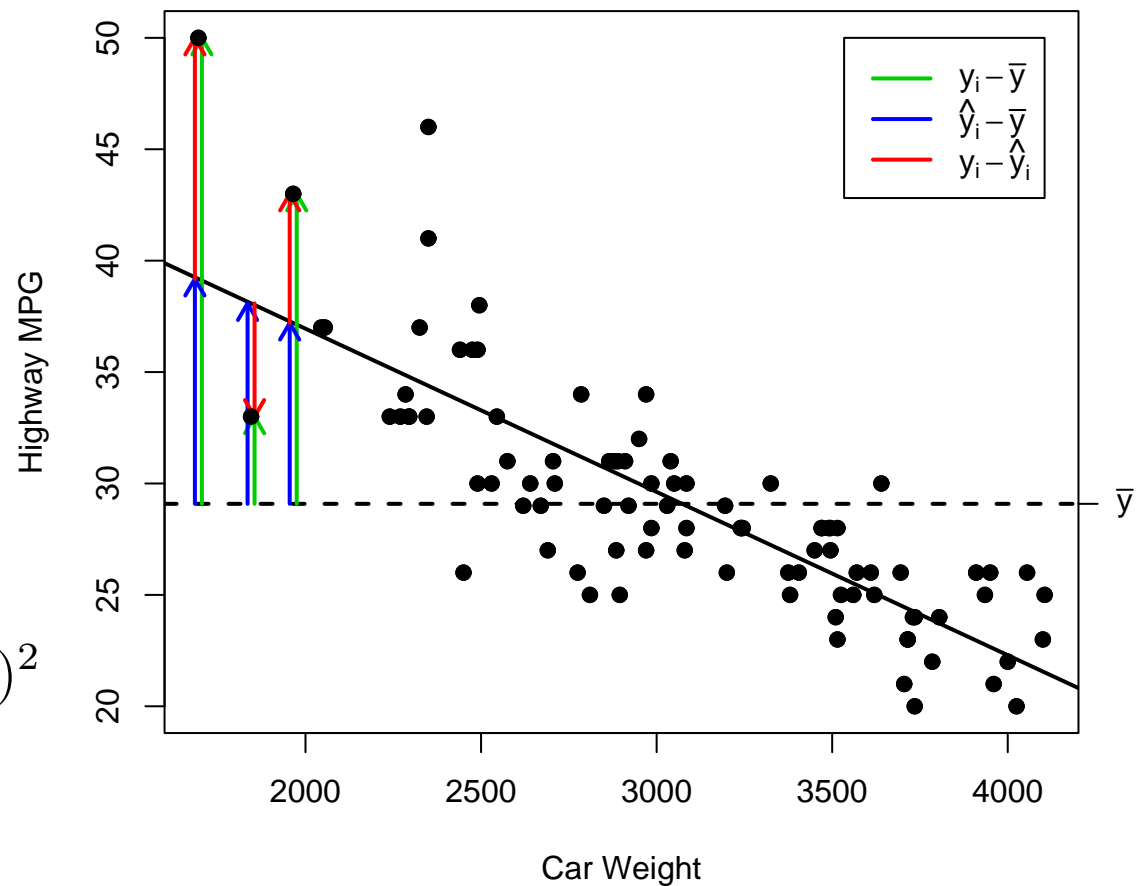
$$y_i - \bar{y}$$

This can be broken down as

$$(y_i - \bar{y}) = (\hat{y}_i - \bar{y}) + (y_i - \hat{y}_i)$$

It is possible to show that

$$\sum (y_i - \bar{y})^2 = \sum (\hat{y}_i - \bar{y})^2 + \sum (y_i - \hat{y}_i)^2$$



We can rewrite this formula as

$$SST = SSM + SSE$$

where

$$SST = \sum (y_i - \bar{y})^2 \quad (\text{Total sums of squares})$$

$$SSM = \sum (y_i - \bar{y})^2 \quad (\text{Model SS})$$

$$SSE = \sum (y_i - \bar{y})^2 \quad (\text{Error or Residual SS})$$

If the slope  $\beta_1 = 0$ , the observations can be viewed as coming from a single population with mean  $\mu_y$  with the variance described by the sample variance

$$s_y^2 = \frac{\sum (y_i - \bar{y})^2}{n - 1} = \frac{SST}{n - 1}$$

You can think of SST as the total error variability of the 0 slope model.

As we have seen before,  $n - 1$  is the degrees of freedom for the single population model and  $n - 1$  is the degrees of freedom for error in the simple linear regression model.

We can breakdown the degrees of freedom like we did the sums of squares

$$DFT = DFM + DFE$$

For simple linear regression

$$DFT = n - 1 \quad (\text{Total degrees of freedom})$$

$$DFM = 1 \quad (\text{Model df})$$

$$DFE = n - 2 \quad (\text{Error or Residual df})$$

Instead of looking at the sums of squares, we can also look at the mean squares (variability per degree of freedom).

$$MS = \frac{\text{sums of squares}}{\text{degrees of freedom}}$$

So

$$MSE = \frac{SSE}{n - 2} = s^2$$
$$MSM = \frac{SSM}{1}$$

We can also fit correlation into this approach to the simple linear regression model. It is possible to show that

$$r^2 = \frac{SSM}{SST}$$

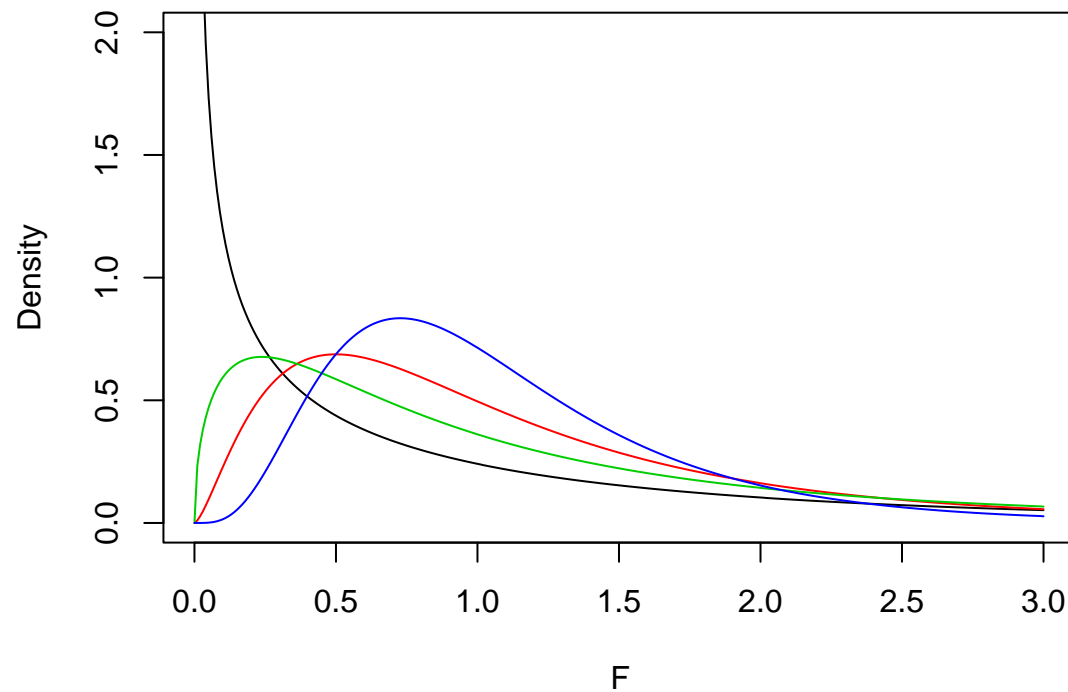
## ANOVA $F$ Test

Instead of using the  $t$  test to investigate the hypotheses  $H_0 : \beta_1 = 0$  vs  $H_A : \beta_1 \neq 0$ , we can look at the ratio

$$F = \frac{MSM}{MSE}$$

If the null hypothesis is false,  $MSE$  should be small and  $MSM$  should be large (leading to  $F > 1$ ). If  $H_0$  is true,  $MSE \approx MSM$  ( $F \approx 1$ ).

The sampling distribution of  $F$  is an  $F$  distribution with 1 and  $n - 2$  degrees of freedom ( $F(1, n - 2)$ )



The  $p$ -value for the  $F$  test is

$$p\text{-value} = P[F(1, n - 2) \geq F_{obs}]$$



## ANOVA Table

Source	DF	SS	MS	F
Model	1	$SSM = \sum(\hat{y}_i - \bar{y})^2$	$MSM = \frac{SSM}{DFM}$	$F = \frac{MSM}{MSE}$
Error	$n - 2$	$SSE = \sum(y_i - \hat{y}_i)^2$	$MSE = \frac{SSE}{DFE}$	
Total	$n - 1$	$SST = \sum(y_i - \bar{y})^2$		

Source	SS	df	MS	Number of obs =	93
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So we have two tests for examining

$$H_0 : \beta_1 = 0 \quad \text{versus} \quad H_A : \beta_1 \neq 0$$

In fact we really only have one, since it's possible to show that  $t^2 = F$  and the  $p$ -values for the two tests are the same.

In the example  $19.35^2 = 374.42$  (within rounding).

The  $F$  test is more useful for multiple regression models and in that situation it looks at more complicated hypotheses.

## Inference for Correlation

There is a third approach to examining whether the data is better described by a line with slope 0.

If there is no correlation between  $x$  and  $y$  ( $\rho = 0$ ), the population regression line will have slope  $\beta_1 = 0$ .

The usual test statistic for examining  $H_0 : \rho = 0$  is

$$t = \frac{r\sqrt{n-2}}{\sqrt{1-r^2}}$$

This statistic has a  $t(n-2)$  distribution.

The assumption behind this sampling distribution is that  $x$  and  $y$  are jointly normally distributed.

Getting  $p$ -values for this test statistic is similar to other  $t$  tests.

$$\begin{array}{ll} H_A : \rho < 0 & p\text{-value} = P[T \leq t_{obs}] \\ H_A : \rho > 0 & p\text{-value} = P[T \geq t_{obs}] \\ H_A : \rho \neq 0 & p\text{-value} = 2 \times P[T \geq |t_{obs}|] \end{array}$$

Should we be confused by have two different  $t$  tests in the linear regression setting? No, as this  $t$  test on correlation is exactly the same as the  $t$  on the slope.

It is possible to show

$$\frac{b_1}{SE_{b_1}} = \frac{r\sqrt{n-2}}{\sqrt{1-r^2}}$$

The following dataset I'll describe in more detail next class, but I want to show how the different approaches all tie in together.

Source	SS	df	MS	Number of obs =	26
Model	4510.59756	1	4510.59756	F( 1, 24) =	0.62
Residual	174203.6	24	7258.48334	Prob > F =	0.4382
Total	178714.198	25	7148.5679	R-squared =	0.0252
				Adj R-squared =	-0.0154
				Root MSE =	85.197

sales	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
Promotion	7.332297	9.301349	0.79	0.438	-11.86474	26.52934
_cons	130.5569	53.00137	2.46	0.021	21.16744	239.9463

. pwcorr sales promotion, sig (Pairwise Correlations)

	sales	promotion
sales	1.0000	
promotion	0.1589	1.0000
	0.4382	

## Theoretical Aside

It is possible to link the parameters from a bivariate normal model to the population regression line model.

This is what motivates the relationship between testing whether a correlation is 0 and whether a slope is 0.

$$\begin{aligned}\beta_1 &= \rho \frac{\sigma_y}{\sigma_x} \\ \beta_0 &= \mu_y - \beta_1 \mu_x \\ \sigma_\epsilon &= \sigma_y \sqrt{1 - \rho^2}\end{aligned}$$