

1. Rice 4.94

$$E\left[\frac{1}{X}\right] = \int_{10}^{20} \frac{1}{10} \frac{1}{x} dx = \frac{\ln 2}{10} = 0.06931472$$

$$E\left[\frac{1}{X^2}\right] = \int_{10}^{20} \frac{1}{10} \frac{1}{x^2} dx = \frac{1}{200} = 0.005$$

$$\text{Var}\left(\frac{1}{X}\right) = 0.005 - 0.06931472^2 = 0.00019751$$

Based on  $E(X) = 15$ ,  $\text{Var}(X) = \frac{25}{3}$ . Let  $f(x) = \frac{1}{x}$ ,  $f'(x) = \frac{-1}{x^2}$ , and  $f''(x) = \frac{2}{x^3}$  so

$$f(x) \approx \frac{1}{E[X]} - \frac{1}{(E[X])^2}(x - E[X])$$

and

$$\text{Approx}\left(E\left[\frac{1}{X}\right]\right) = \frac{1}{E[X]} = \frac{1}{15} = 0.06667$$

$$\text{Approx}\left(\text{Var}\left(\frac{1}{X}\right)\right) = \frac{1}{(E[X])^4} \text{Var}(X) = \frac{1}{15^4} \frac{25}{3} = 0.000165$$

The second order approximation to the mean is

$$\text{Approx}\left(E\left[\frac{1}{X}\right]\right) = \frac{1}{E[X]} + \frac{2}{(E[X])^3} \text{Var}(X) = \frac{1}{15} + \frac{1}{15^3} \frac{25}{3} = 0.06914$$

2. Rice 4.98

(a) Let  $E[R] = \mu_R$ ,  $\text{Var}(R) = \sigma_R^2$ ,  $E[\theta] = \mu_\theta$ ,  $\text{Var}(\theta) = \sigma_\theta^2$ . Then for  $g(R, \theta) = R \sin \theta$

$$Y \approx \mu_R \sin \mu_\theta + \sin \mu_\theta (R - \mu_R) + R \cos \mu_\theta (\theta - \mu_\theta)$$

$$\text{Approx}(\text{Var}(Y)) = \sigma_R^2 \sin^2 \mu_\theta + \mu_R^2 \sigma_\theta^2 \cos^2 \mu_\theta$$

(b) if known  $R = r$ , then  $Y = r \sin \theta$  and

$$\text{Var}(Y) \approx r^2 \cos^2 \theta \text{Var}(\theta)$$

$\theta = 0$  leads to the most variable estimated altitude as

$$\frac{d}{d\theta} r^2 \sigma_\theta^2 \cos^2 \theta = 2r^2 \sigma_\theta^2 \sin \theta \cos \theta$$

takes the value 0 for  $\theta = 0$  and  $\frac{\pi}{2}$ . By checking the function,  $\theta = 0$  corresponds to the maximum and  $\frac{\pi}{2}$  to a minimum.

This answer is also supported by noting that

$$\frac{dY}{d\theta} = r \cos \theta$$

is maximized by  $\theta = 0$ .

### 3. Rice 5.5

Since  $X_n \sim \text{Bin}(n, p)$ ,

$$M_{X_n}(t) = (1 + p(e^t - 1))^n$$

Making the substitution  $p = \frac{\lambda}{n}$

$$M_{X_n}(t) = \left(1 + \frac{\lambda(e^t - 1)}{n}\right)^n$$

Then as  $(1 + \frac{a}{n})^n \rightarrow e^a$ ,

$$M_{X_n}(t) \rightarrow \exp(\lambda(e^t - 1))$$

which is the moment generating function of a  $\text{Pois}(\lambda)$  random variable, implying  $X_n \xrightarrow{\mathcal{D}} \text{Pois}(\lambda)$ .

### 4. Rice 5.12

Let

$$X_i \sim \text{Unif}(-0.5, 0.5) \quad E[X_i] = 0 \quad \text{Var}(X_i) = \frac{1}{12}$$

So

$$\frac{\sum_{i=1}^{100} X_i - 0}{\sqrt{\frac{1}{12} \cdot 100}} \underset{\text{approx.}}{\sim} N(0, 1)$$

implying

$$\begin{aligned} P \left[ \left| \sum_{i=1}^{100} X_i \right| > 1 \right] &\approx P \left[ \left| \frac{\sum_{i=1}^{100} X_i}{\sqrt{\frac{100}{12}}} \right| > \frac{\sqrt{3}}{5} \right] \\ &= 2 \left( 1 - \Phi \left( \frac{\sqrt{3}}{5} \right) \right) = 0.73 \end{aligned}$$

Similarly

$$P \left[ \left| \sum_{i=1}^{100} X_i \right| > 2 \right] \approx 2 \left( 1 - \Phi \left( 2 \frac{\sqrt{3}}{5} \right) \right) = 0.4884224$$

and

$$P \left[ \left| \sum_{i=1}^{100} X_i \right| > 5 \right] \approx 2 \left( 1 - \Phi \left( \sqrt{3} \right) \right) = 0.08326452$$

5. Rice 5.14

The result of step  $i$  follows the distribution  $P[X_i = +50] = \frac{2}{3}$  (North) and  $P[X_i = -50] = \frac{1}{3}$  (South). Then

$$\begin{aligned} E[X_i] &= 50 \frac{2}{3} - 50 \frac{1}{3} = \frac{50}{3} \\ \text{Var}(X_i) &= 50^2 \frac{2}{3} + (-50)^2 \frac{1}{3} - \left( \frac{50}{3} \right)^2 \\ &= \frac{2500}{3} - \frac{2500}{9} = \frac{5000}{9} \end{aligned}$$

The final position is given by  $S_{60} = \sum_{i=1}^{60} X_i$ . Then

$$\begin{aligned} E[S_{60}] &= 60 \frac{50}{3} = 1000 \\ \text{Var}(S_{60}) &= 60 \frac{5000}{9} = 33333.3 \end{aligned}$$

So  $S_{60} \overset{\text{approx.}}{\sim} N(1000, 33333.3)$ . So the most likely position is 1000cm (= 10 m) to the north of his starting position. The standard deviation of the difference from this is 182.6 cm.

6. Rice 5.23

Let  $Z_1, \dots, Z_n$  be a set of iid indicator variables. Then  $P[Z_i = 1] = A$  and  $P[Z_i = 0] = 1 - A$ , so  $E[Z_i] = A$ . Then by the weak law of large numbers,  $\bar{Z} \xrightarrow{P} E[Z_i] = A$ .

7. Rice 5.24

Note that  $\sum_{i=1}^n Z_i \sim \text{Bin}(n, A)$ . So as  $n \rightarrow \infty$ ,  $\hat{A} \overset{\text{approx.}}{\sim} N\left(A, \frac{A(1-A)}{n}\right)$ . So we need to choose  $n$  such that

$$\begin{aligned} 0.99 &= P[|\hat{A} - A| < 0.01] \\ &= P\left[\left|\frac{(\hat{A} - A)\sqrt{n}}{\sqrt{A(1-A)}}\right| < \frac{0.01\sqrt{n}}{\sqrt{A(1-A)}}\right] \\ &\approx P\left[|Z| < \frac{0.01\sqrt{n}}{\sqrt{A(1-A)}}\right] \end{aligned}$$

This implies that (with  $A = 0.2$ )

$$\frac{0.01\sqrt{n}}{\sqrt{0.16}} = 2.576$$

or

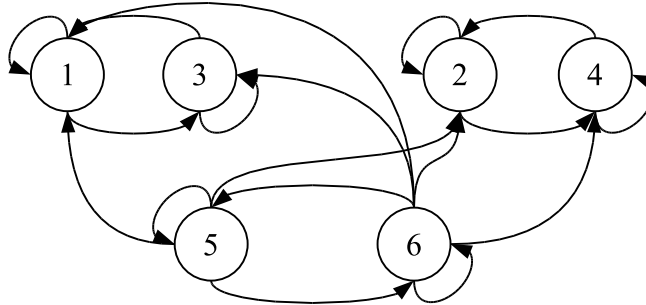
$$n = \frac{2.576^2 \cdot 0.16}{0.01^2} = 10617.2$$

8.  $nF_n(x) \sim \text{Bin}(n, F(x))$ , so by the weak law of law numbers,  $F_n(x)$  converges in probability its success probability  $F(x)$ .

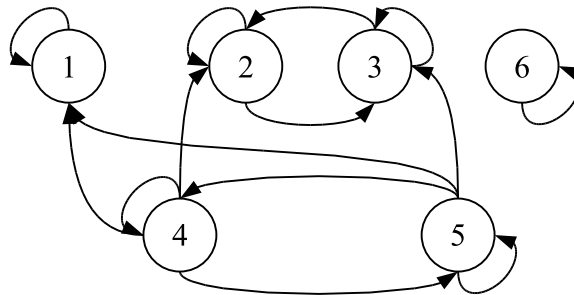
Another approach, is to note that  $E[F_n(x)] = F(x)$  and  $\text{Var}(F_n(x)) = \frac{F(x)(1-F(x))}{n}$ . By Chebyshev's inequality,

$$P[|F_n(x) - F(x)| \geq \delta] \leq \frac{F(x)(1 - F(x))}{n\delta^2} \rightarrow 0$$

9. (a) (1,3), (2,4), (5,6) are communicating classes, with (1,3), (2,4) being recurrent, and (5,6) being transient.



- (b) (1), (2,3),(6) are communicating recurrent classes, and (4,5) is a communicating transient class.



10. Let states 1,2,3 are notations for states C, S, G respectively. Then the transition matrix

$$P = \begin{bmatrix} 0.7 & 0.2 & 0.1 \\ 0.4 & 0.3 & 0.3 \\ 0.2 & 0.4 & 0.4 \end{bmatrix}$$

To find stationary distribution  $\boldsymbol{\pi} = (\pi_1, \pi_2, \pi_3)$  is to solve

$$\boldsymbol{\pi}P = \boldsymbol{\pi}$$

subject to  $\pi_1 + \pi_2 + \pi_3 = 1$ . The solution is

$$\boldsymbol{\pi} = (\pi_1, \pi_2, \pi_3) = \left( \frac{30}{59}, \frac{16}{59}, \frac{13}{59} \right)$$

The the proportion of days that Buffy cheerful is  $\pi_1 = \frac{30}{59}$  and the long-run average number of iterations between gloomy days is  $\frac{1}{\pi_3} = \frac{59}{13} = 4.54$  day by applying the result that

$$\pi_3 = \frac{1}{E_3(T_y)}$$

11. There are three possible states in this problem: state 1 is that both switches are off, state 2 is that one is off and the other is on, and state 3 is that both switches are on. Then the transition probability matrix is

$$P = \begin{bmatrix} \frac{3}{4} \times \frac{3}{4} & 2\frac{3}{4} \times \frac{1}{4} & \frac{1}{4} \times \frac{1}{4} \\ \frac{1}{2} \times \frac{1}{2} & 2\frac{1}{2} \times \frac{1}{2} & \frac{1}{2} \times \frac{1}{2} \\ \frac{1}{4} \times \frac{1}{4} & 2\frac{3}{4} \times \frac{1}{4} & \frac{3}{4} \times \frac{3}{4} \end{bmatrix} = \begin{bmatrix} \frac{9}{16} & \frac{6}{16} & \frac{1}{16} \\ \frac{4}{16} & \frac{8}{16} & \frac{4}{16} \\ \frac{1}{16} & \frac{6}{16} & \frac{9}{16} \end{bmatrix}$$

Solving the equations

$$\boldsymbol{\pi}P = \boldsymbol{\pi}$$

with  $\pi_1 + \pi_2 + \pi_3 = 1$ , then we get

$$\boldsymbol{\pi} = (\pi_1, \pi_2, \pi_3) = \left( \frac{2}{7}, \frac{3}{7}, \frac{2}{7} \right)$$

So the fraction of time that both switches are on is  $\frac{2}{7}$ .

### Suggested Problems

1. Rice 5.13

The result of step  $i$  follows the distribution  $P[X_i = +50] = \frac{1}{2}$  (North) and  $P[X_i = -50] = \frac{1}{2}$  (South). Then

$$E[X_i] = 50\frac{1}{2} - 50\frac{1}{2} = 0$$

$$\text{Var}(X_i) = 50^2\frac{1}{2} + (-50)^2\frac{1}{2} - 0 = 2500$$

The final position is given by  $S_{60} = \sum_{i=1}^{60} X_i$ . Then

$$E[S_{60}] = 60 \times 0 = 0$$

$$\text{Var}(S_{60}) = 60 \times 2500 = 150,000$$

So  $S_{60} \overset{\text{approx.}}{\sim} N(0, 150000)$ . So the most likely position is the initial starting position. The standard deviation of the difference from this is 387.3 cm.

2. Rice 5.17

$$P[|\bar{X} - \mu| < 1] = P\left[\left|\frac{\sqrt{n}(\bar{X} - \mu)}{5}\right| < \frac{\sqrt{n}}{5}\right] \\ \approx P\left[|Z| < \frac{\sqrt{n}}{5}\right] = 0.95$$

So need  $n$  such that

$$\frac{\sqrt{n}}{5} = 1.96 \quad \Rightarrow \quad n = (5 \times 1.96)^2 = 96.04$$

The Chebyshev approach gives

$$P[|\bar{X} - \mu| < 1] \geq 1 - \frac{\text{Var}(\bar{X})}{1^2} = 1 - \text{Var}(\bar{X})$$

Thus we need to solve

$$\text{Var}(\bar{X}) = \frac{25}{n} = 0.05$$

which yields  $n = \frac{25}{0.05} = 500$  as a upper bound for  $n$ .

3. Rice 5.26

$$S \sim \text{Bin}(25, 0.3) \approx N(7.5, 5.25)$$

Let  $Y \sim N(7.5, 5.25)$

| x  | Exact<br>$P[S \leq x]$ | Normal Approximation<br>$P[Y \leq x]$ | Corrected Approximation<br>$P[Y \leq x + 0.5]$ |
|----|------------------------|---------------------------------------|--|
| 5  | 0.1935                 | 0.1376                                | 0.1913   |
| 7  | 0.5118                 | 0.4136                                | 0.5000   |
| 9  | 0.8106                 | 0.7437                                | 0.8086   |
| 11 | 0.9558                 | 0.9367                                | 0.9596   |