

General Linear Statistical Models - Part III

Statistics 135

Autumn 2005



Interaction Models

Lets examine two models involving `Weight` and `Domestic` in the `cars93` dataset.

```
weight.domestic.lm <- lm(HighFuel ~ Weight + Domestic,  
                        data=cars93)
```

```
weight.domestic.int.lm <- lm(HighFuel ~ Weight * Domestic,  
                             data=cars93)
```

Remember that the second model is a shorthand for

```
HighFuel ~ Weight + Domestic + Weight : Domestic
```

```
> summary(weight.domestic.lm)
```

```
Residuals:
```

Min	1Q	Median	3Q	Max
-0.781506	-0.244967	0.002068	0.180682	0.922104

```
Coefficients:
```

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	9.923e-01	1.853e-01	5.355	6.5e-07	***
Weight	8.354e-04	6.065e-05	13.774	< 2e-16	***
DomesticDomestic	-3.449e-02	7.120e-02	-0.484	0.629	

```
> summary(weight.domestic.int.lm)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.78647	-0.21346	-0.03952	0.17163	0.99145

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	6.264e-01	2.504e-01	2.501	0.0142	*
Weight	9.597e-04	8.347e-05	11.498	<2e-16	***
DomesticDomestic	7.421e-01	3.721e-01	1.994	0.0492	*
Weight:DomesticDomestic	-2.529e-04	1.190e-04	-2.125	0.0364	*

Both models fit give regression lines for HighFuel vs Weight. The first is the additive model, which is of the form

$$y_i = \beta_0 + \beta_1 w_i + \beta_2 d_i + \epsilon_i$$

The second model is of the form

$$y_i = \beta_0 + \beta_1 w_i + \beta_2 d_i + \beta_3 w_i d_i + \epsilon_i$$

where d_i is 1 for domestic cars and 0 for foreign cars

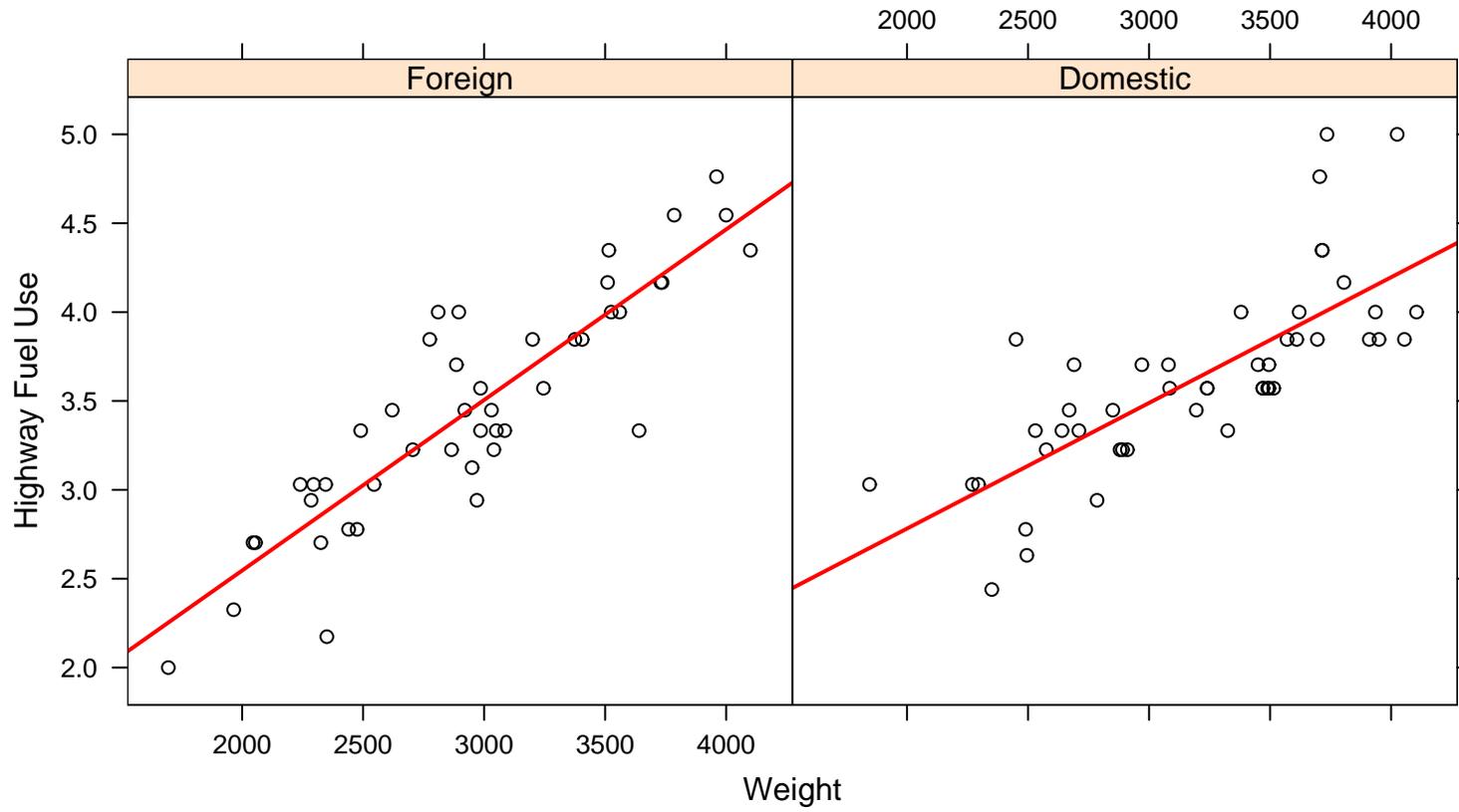
These can be written as

$$y_i = \begin{cases} \beta_0 + \beta_1 w_i + \epsilon_i & \text{Foreign car} \\ \underbrace{(\beta_0 + \beta_2)}_{\beta_0^*} + \beta_1 w_i + \epsilon_i & \text{Domestic car} \end{cases}$$

and

$$y_i = \begin{cases} \beta_0 + \beta_1 w_i + \epsilon_i & \text{Foreign car} \\ \underbrace{(\beta_0 + \beta_2)}_{\beta_0^*} + \underbrace{(\beta_1 + \beta_3)}_{\beta_1^*} w_i + \epsilon_i & \text{Domestic car} \end{cases}$$

The second model is an example of an interaction. In this case, the effect of weight depends on whether the car is domestically made or not. It is fitting the equivalent to what is displayed in the figure



As mentioned last class to indicate interactions in **S**, you use either * or :. Usually you want to use *, as its shorter and it also leads to good statistical practice. Consider the models (where x and y are quantitative)

$$z \sim x*y$$

$$z \sim x:y$$

The first model is equivalent to $z \sim x + y + x:y$, a standard model of interest. The second model would fit

$$z_i = \beta_0 + \beta_1 x_i y_i$$

Usually you don't want to fit a model with the lower order terms missing. For example, with `HighFuel ~ Weight:Domestic`, **S** would be fitting

$$y_i = \begin{cases} \beta_0 + \beta_1 w_i + \epsilon_i & \text{Foreign car} \\ \beta_0 + \epsilon_i & \text{Domestic car} \end{cases}$$

Not a particularly reasonable model.

Note: this idea is a good rule of thumb. There may be situations where you might want to include higher order interactions but drop out lower order ones.

One possibility of this is

$$y_i = \beta_1 x_i + \beta_2 x_i d_i + \epsilon_i$$

Here the main effect for d_i is missing. This model is describing regression through the origin for x_i , with different slopes for different levels of d_i .

If terms get repeated in a model description, the repeats get dropped, so don't worry about them. So for example

$$y \sim A*B + B*C$$

is a fine way of describing the model

$$y \sim A + B + C + A:B + B:C$$

Polynomial Models

It is easy to fit polynomial models in **S**. However there is a slight trick to it (particularly in **R**). One approach you might consider is a call like

```
> weight2a.lm <- lm(HighFuel ~ Weight + Weight^2, data=cars93)
> summary(weight2a.lm)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.768280	-0.239028	0.005072	0.199289	0.909606

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	9.940e-01	1.845e-01	5.387	5.56e-07	***
Weight	8.290e-04	5.898e-05	14.057	< 2e-16	***

This only fits the linear term (the Weight^2 term gets dropped). (Actually this only happens in **R**. It will do what you expect in **S-Plus**.)

Instead you need to use the `I()` operator as in the following

```
> weight2.lm <- lm(HighFuel ~ Weight + I(Weight^2), data=cars93)
> summary(weight2.lm)
```

Call:

```
lm(formula = HighFuel ~ Weight + I(Weight^2), data = cars93)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.76605	-0.23896	0.01345	0.19332	0.91241

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	8.278e-01	8.696e-01	0.952	0.344
Weight	9.430e-04	5.855e-04	1.611	0.111
I(Weight^2)	-1.879e-08	9.607e-08	-0.196	0.845

Residual standard error: 0.3355 on 90 degrees of freedom
Multiple R-Squared: 0.6848, Adjusted R-squared: 0.6778
F-statistic: 97.78 on 2 and 90 DF, p-value: < 2.2e-16

```
> anova(weight2.lm)
Analysis of Variance Table
```

Response: HighFuel

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
Weight	1	22.0027	22.0027	195.5175	<2e-16	***
I(Weight^2)	1	0.0043	0.0043	0.0383	0.8454	
Residuals	90	10.1282	0.1125			

This gives you what you want.

Removing Terms from Models

It is also possible to remove terms from models. For example

$$y \sim A + B + C + A:B + A:B + B:C$$

could have been written

$$y \sim A*B*C - A:B:C$$

so it can be used as a shorthand to write more complicated models.

Another situation where it is useful to to compare two models. Consider in the crab example where we want to compare the models

$$RW \sim \text{sex} * \text{sp}$$

and

$$RW \sim \text{sex} + \text{sp}$$

One way of doing this is by

```
crab.int.lm <- lm(RW ~ sex * sp, data=crabs)
crab.add.lm <- update(crab.int.lm, . ~ . - sex:sp)
```

This could also be done by

```
crab.add2.lm <- lm(RW ~ sex + sp, data=crabs)
crab.int2.lm <- update(crab.add2.lm, . ~ . + sex:sp)
```

To see whether the interaction model gives a better fit, we can look at the command

```
> anova(crab.add.lm, crab.int.lm)
```

Analysis of Variance Table

Model 1: RW ~ sex + sp

Model 2: RW ~ sex * sp

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	197	1074.4				
2	196	1016.4	1	58.0	11.184	0.0009884 ***

Note that this isn't needed for this example as

```
> anova(crab.int.lm)
Analysis of Variance Table
```

```
Response: RW
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
sex	1	112.05	112.05	21.608	6.133e-06	***
sp	1	131.38	131.38	25.336	1.087e-06	***
sex:sp	1	58.00	58.00	11.184	0.0009884	***
Residuals	196	1016.36	5.19			

gives the same information.

Another useful situation where removing a term may be useful is to get rid of the intercept. For example to fit a regression through the origin you can do

```
> weight.orig.lm <- lm(HighFuel ~ Weight - 1, data=cars93)
> summary(weight.orig.lm)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.820327	-0.227628	-0.009304	0.320788	1.050421

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
Weight	1.141e-03	1.263e-05	90.33	<2e-16 ***

Residual standard error: 0.3811 on 92 degrees of freedom

Multiple R-Squared: 0.9888, Adjusted R-squared: 0.9887

F-statistic: 8159 on 1 and 92 DF, p-value: < 2.2e-16

There is some evidence for this model. First the physics suggests it. In addition

```
> weight.lm <- lm(HighFuel ~ Weight, data=cars93)
> summary(weight.lm)
```

Call:

```
lm(formula = HighFuel ~ Weight, data = cars93)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.768280	-0.239028	0.005072	0.199289	0.909606

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	9.940e-01	1.845e-01	5.387	5.56e-07	***
Weight	8.290e-04	5.898e-05	14.057	< 2e-16	***

Residual standard error: 0.3337 on 91 degrees of freedom

Multiple R-Squared: 0.6847, Adjusted R-squared: 0.6812

F-statistic: 197.6 on 1 and 91 DF, p-value: < 2.2e-16

```
> anova(weight.orig.lm, weight.lm)
```

Analysis of Variance Table

Model 1: HighFuel ~ Weight - 1

Model 2: HighFuel ~ Weight

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	92	13.3643				
2	91	10.1325	1	3.2318	29.025	5.564e-07 ***

Removing the intercept is useful in some ANOVA models as it gives another parameterization. For example,

```
> type.lm
```

Call:

```
lm(formula = HighFuel ~ Type, data = cars93)
```

Coefficients:

(Intercept)	TypeLarge	TypeMidsize	TypeSmall	TypeSporty	TypeVan
3.3768	0.3725	0.3965	-0.4979	0.1475	1.2098

```
> type.noint.lm
```

Call:

```
lm(formula = HighFuel ~ Type - 1, data = cars93)
```

Coefficients:

TypeCompact	TypeLarge	TypeMidsize	TypeSmall	TypeSporty	TypeVan
3.377	3.749	3.773	2.879	3.524	4.587

In the second approach, the parameter estimates are the sample means for each Type.

Another example is

```
> weight.domestic.int.lm
```

```
Call: lm(formula = HighFuel ~ Weight * Domestic, data = cars93)
```

```
Coefficients:
```

(Intercept)	Weight	DomesticDomestic
0.6263581	0.0009597	0.7420544
Weight:DomesticDomestic		
-0.0002529		

```
> weight.domestic.int2.lm
```

```
Call: lm(formula = HighFuel ~ Domestic/Weight - 1, data = cars93)
```

```
Coefficients:
```

DomesticForeign	DomesticDomestic	DomesticForeign:Weight
0.6263581	1.3684125	0.0009597
DomesticDomestic:Weight		
0.0007069		

The first gives the difference in the intercepts and slopes for domestic cars from foreign cars where the second gives the slopes and intercepts for both types.

The / is another way of describing interactions. The for is a / x , where a is a factor and x could be numeric, a factor, or a combination of things. This model says fit the model described by x for each level of a . The specification $a/x - 1$ is equivalent to

$$a + a:x - 1$$

in terms of parameterization.

Prediction

It is easy to make predictions for new or hypothesized observations with the `predict` function. The form of the function is `predict(fit, newdata)`, where `fit` is the result of the `lm` command and `newdata` is a dataframe including all of the variables used in the fitting model

```
> newdata
  Weight Domestic
1   2000 Foreign
2   3000 Domestic
3   4000 Foreign
4   2000 Domestic
5   3000 Foreign
6   4000 Domestic
> predict(weight.domestic.int.lm,newdata)
      1      2      3      4      5      6
2.545835 3.489005 4.465312 2.782141 3.505573 4.195869
```

And to exhibit that different parametrizations give the same fitted values

```
> predict(weight.domestic.int.lm,newdata)
      1      2      3      4      5      6
2.545835 3.489005 4.465312 2.782141 3.505573 4.195869
```

```
> predict(weight.domestic.int2.lm,newdata)
      1      2      3      4      5      6
2.545835 3.489005 4.465312 2.782141 3.505573 4.195869
```